BS 2846-6: 1976 ISO 3301:1975

Guide to

# Statistical interpretation of data —

Part 6: Comparison of two means in the case of paired observations

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The Government department and scientific and industrial organizations marked with an asterisk in the above list were directly represented on the committee entrusted with the preparation of this British Standard.

This British Standard, having been prepared under the direction of The Advisory Committee on Statistical Methods was published under the authority of the Executive Board on 30 June 1976

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# Foreword

The correct interpretation and presentation of test results have been assuming increasing importance in the analysis of data obtained from manufacturing processes based on sample determinations and prototype evaluations in industry, commerce and educational institutions. It was for this reason that Subcommittee 2 of Technical Committee 69, "Applications of Statistical Methods", of the International Organization for Standardization (ISO), was charged with the task of preparing a guide to statistical methods for the interpretation of test results. As international agreement is reached on the statistical tests relevant to specific situations it is proposed to publish them as parts of a revised BS 2846 1957: "The reduction and presentation of experimental results" as follows.

Statistical interpretation of data

- Part 1: Routine analysis of quantitative data;
- Part 2: Estimation of the mean-confidence interval [ISO 2602];
- Part 3: Determination of a statistical tolerance interval [ISO 3207];

- Part 4: Techniques of estimation and tests relating to means and variances [ISO 2854]<sup>1</sup>;

— Part 5: Efficiency of tests relating to means and variances [ISO 3494]<sup>1)</sup>;

— Part 6: Comparison of two means in the case of paired observations [ISO 3301].

A situation often encountered is one in which a decision in favour of one of two alternatives, whether they be manufacturing processes, machines or medical drugs etc., needs to be taken. There exists a number of criteria for making such judgements depending upon whether it is the "average effect" or the number of "individual items being processed that are better than some tolerance" that is of prime importance. This Part of the standard specifies a technique for testing for (significant) differences between two mean values (e.g. two process averages) and consequently is concerned with problems of the former kind. The latter problem is dealt with in Part 3, and also BS 6002

"Sampling procedures and charts for inspection by variables for percent defective".<sup>1)</sup> The technique of this Part provides a procedure for deciding between two alternatives on the basis of differences between their average performances.

This Part of this British Standard is identical with ISO 3301 "Statistical interpretation of data — Comparison of two means in the case of paired observations".

For the purposes of this British Standard the text of ISO 3301 given in this publication should be modified as follows.

**Terminology.** The words "British Standard" should replace "International Standard" wherever they appear.

The decimal point should replace the decimal comma, wherever it appears.

**Cross-reference.** The references to other International Standards should be replaced by references to British Standards as follows.

#### **Reference to ISO Standard**

ISO 2854 "Statistical interpretation of data — Techniques of estimation and tests relating to the means and variances"

#### **Appropriate British Standard**

BS 2846 "Guide to statistical interpretation of data" Part 4 "Techniques of estimation and tests relating to means and variances"<sup>a</sup>

<sup>a</sup> In course of preparation.

 $<sup>^{1)}</sup>$  In course of preparation.

The following British Standards also provide practical guidance on the application of statistical methods.

BS 600, Application of statistical methods to industrial standardization and quality control.

BS 1313, Fraction-defective charts for quality control.

BS 2564, Control chart technique when manufacturing to a specification, with special reference to articles machined to dimensional tolerances.

BS 6000, Guide to the use of BS 6001. Sampling procedures and tables for inspection by attributes.

BS 6001, Sampling procedures and tables for inspection by attributes.

BS 6002<sup>2)</sup>, Sampling procedures and charts for inspection by variables for percent defective.

BS  $\dots^{2)}$ , Determination of repeatability and reproducibility for a standard test method.

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#### Summary of pages

This document comprises a front cover, an inside front cover, pages i to iv, pages 1 to 6 and a back cover.

This standard has been updated (see copyright date) and may have had amendments incorporated. This will be indicated in the amendment table on the inside front cover.

 $<sup>^{2)}\,\</sup>mathrm{In}$  course of preparation.

# **0** Introduction

The method specified in this International Standard, known as the method of paired observations, is a special case of the method described in Table A' of ISO 2854, *Statistical interpretation of data* — *Techniques of estimation and tests relating to means and variances.*<sup>3)</sup>

This special case is mentioned in Section 2 of ISO 2854 immediately after the numerical illustration of Table A', and a complete example of applications of the method of paired comparisons has been given in Annex A of that International Standard. The importance and wide applicability of the method justify a separate International Standard being devoted to it.

## 1 Scope

This International Standard specifies a method for comparing the mean of a population of differences between paired observations with zero or any other preassigned value.

## **2 Definition**

#### paired observations

two observations  $x_i$  and  $y_i$  of a certain property or characteristic are said to be paired if they are made:

— on the same element i from a population of elements but under different conditions (for example, comparison of results of two methods of analysis on the same product);

— on two distinct elements considered similar in all respects except for the systematic difference which is the subject of the test (for example, comparison of the yield from adjacent plots sown with two distinct varieties of seed).

however, it should be noted that in the second case the efficiency of the test depends on the validity of the hypothesis that there is no other systematic difference between the individuals in the same pair other than the systematic difference under test

# **3 Field of application**

The method may be applied to establish a difference between two treatments. In this case, the observations x; are carried out after the first treatment and  $y_i$  after the second treatment. The two series of results of the observations are not independent because each result  $x_i$  of the first series (first treatment) is associated with a result y, of the second series (second treatment). The term "treatment" should be understood in a wide sense. The two treatments to be compared may, for instance, be two test methods, two measuring instruments or two laboratories, in order to detect a possible systematic error. Two treatments carried out successively on the same experimental material might interact and the value obtained might depend on the order. Good experimental design should enable such biases to be eliminated. Alternatively, only one treatment may be applied and its effect may be compared to the absence of treatment; the purpose of this comparison is then to establish the effect of that treatment.

# 4 Conditions for application

The method can be applied validly if the following two conditions are satisfied:

— the series of differences  $d_i = x_i - y_i$  can be considered as a series of independent random items;

— the distribution of the differences  $d_i = x_i - y_i$ between the paired observations is supposed to be normal or approximately normal.

If the distribution of these differences deviates from the normal, the technique described remains valid, provided the sample size is sufficiently large; the greater the deviation from normality, the larger the sample size required. Even in extreme cases, however, a sample size of 100 will be sufficient for most practical applications.

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<sup>&</sup>lt;sup>3)</sup> At present at the stage of draft.

# **5** Formal presentation of calculations

Problem studied			
Experimental conditions			
Statistical data		Calculations	
Sample size:			
<i>n</i> =		$\overline{d} = \frac{1}{n} (\Sigma x_i - \Sigma y_i)$	
Sum of the observed va	lues:		
$\Sigma x_i =$	$\Sigma y_i =$	$=\frac{1}{n}\Sigma d_i =$	
Sum of the differences:			
$\Sigma d_i =$		$s_d^2 = \frac{1}{n-1} \Big[ \Sigma d_i^2 - \frac{1}{n} (\Sigma d_i)^2 \Big] =$	
Sum of the squares of t	he differences:		
$\Sigma d_i^2 =$		$\mathbf{c}_d^\star = \sqrt{s_d^2} =$	
Given value:			
<i>d</i> <sub>o</sub> =		$A_1 = [t_{1-\alpha}(v)/\sqrt{n}]\sigma_d^* =$	
Degrees of freedom:			
v = n - 1 =		$A_2 = [t_{1-\alpha/2}(v)/\sqrt{n}]\sigma_d^* =$	
Chosen significance level:			
α =			

## Results

Two-sided case:

The hypothesis that the population mean of the difference is equal to  $d_0$  (null hypothesis) is rejected if:

$$\left| \overline{d} - d_0 \right| > A_2$$

One-sided cases:

a) The hypothesis that the population mean of the differences is at least equal to  $d_0$  (null hypothesis) is rejected if:

 $\overline{d} < d_0 - A_1$ 

b) The hypothesis that the population mean of the differences is at most equal to  $d_0$  (null hypothesis) is rejected if:

$$\overline{d} > d_0 + A_1$$

NOTE  $t_{1-\alpha}(v)$  is the fractile of order  $1-\alpha$  of Student's variate *t* with *v* degrees of freedom. The values of  $t_{1-\alpha}(v)/\sqrt{n}$  are given in Table 1.

	Two-sided case		On	One-sided case	
v = n - 1	$\frac{t_{0, 975}}{\sqrt{n}}$	$\frac{t_{0, 995}}{\sqrt{n}}$	$\frac{t_{0, 95}}{\sqrt{n}}$	$\frac{t_{0,99}}{\sqrt{n}}$	
1	8,985	45,013	4,465	22,501	
2	2,434	5,730	1,686	4,021	
3	1,591	2,920	1,177	2,270	
4	1,242	2,059	0,953	1,676	
5	1,049	1,646	0,823	1,374	
6	0,925	1,401	0,734	1,188	
7	0,836	1,237	0,670	1,060	
8	0,769	1,118	0,620	0,966	
9	0,715	1,028	0,580	0,892	
10	0,672	0,956	0,546	0,833	
11	0,635	0,897	0,518	0,785	
12	0,604	0,847	0,494	0,744	
13	0,577	0,805	0,473	0,708	
14	0,554	0,769	0,455	0,678	
15	0,533	0,737	0,438	0,651	
16	0,514	0,708	0,423	0,626	
17	0,497	0,683	0,410	0,605	
18	0,482	0,660	0,398	0,586	
19	0,468	0,640	0,387	0,568	
20	0,455	0,621	0,376	0,552	
21	0,443	0,604	0,367	0,537	
22	0,432	0,588	0,358	0,523	
23	0,422	0,573	0,350	0,510	
24	0,413	0,559	0,342	0,498	
25	0,404	0,547	0,335	0,487	
26	0,396	0,535	0,328	0,477	
27	0,388	0,524	0,322	0,467	
28	0,380	0,513	0,316	0,458	
29	0,373	0,503	0,310	0,449	
30	0,367	0,494	0,305	0,441	
40	0,316	0,422	0,263	0,378	
50	0,281	0,375	0,235	0,337	
60	0,256	0,341	0,214	0,306	
70	0,237	0,314	0,198	0,283	
80	0,221	0,293	0,185	0,264	
90	0,208	0,276	0,174	0,248	
100	0,197	0,261	0,165	0,235	
200	0,139	0,183	0,117	0,165	
500	0,088	0,116	0,074	0,104	
$\infty$	0	0	0	0	

**Table 1** — Values of the ratio  $t_{1-\alpha}(v)/\sqrt{n}$  for v = n-1

**Example:** The data tabled below were collected during an investigation designed to determine whether the average rate of shaft-wear caused by various bearing metals in an internal combustion engine differed between metals.

Table 2 —	Shaft-wear	after a give	n working time	e in 0.000 01 in
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Shaft	Wear	Difference	
i	copper-lead $x_i$	white metal $y_i$	$d_i = x_i - y_i$
1	3.5	1.5	2.0
2	2.0	1.3	0.7
3	4.7	4.5	0.2
4	2.8	2.5	0.3
5	6.5	4.5	2.0
6	2.2	1.7	0.5
7	2.5	1.8	0.7
8	5.8	3.3	2.5
9	4.2	2.3	1.9
Total	34.2	23.4	10.8

Technical characteristics		
Statistical data	Calculations	
Sample size:		
n = 9	$\overline{d} = \frac{1}{9}(34.2 - 23.4) = 1.2$	
Sum of the observed values:		
$\Sigma x_i = 34.2 \qquad \qquad \Sigma y_i = 23.4$	$s_d^2 = \frac{1}{8}19.22 - \frac{10.8^2}{9} = 0.782$ 5	
Sum of the differences:		
$\Sigma d_i = 10.8$	$\sigma_d^* = \sqrt{0.7825} = 0.8846$	
Sum of the squares of the differences:		
$\Sigma d_i^2 = 19.22$	$t_{0.995}/\sqrt{9} = 1.118$	
Given value:		
$d_{\rm o} = 0$	$A_2 = 1.118 \times 0.884 \ 6 = 0.99$	
Degrees of freedom:		
v = 8		
Chosen significance level:		
$\alpha = 0.01$		
Result		
Comparison of the population mean with the given	value 0:	
Two-sided case:		
$\left  \overline{d} - d_0 \right  = 1.2 > 0.99$		

The hypothesis of the equality of the rate of shaft-wear by the two metals is rejected at the 1 % level.

## 6 Errors of the second kind

The probability of rejecting the null hypothesis when it is true is at most equal to the significance level  $\alpha$ . Rejecting the null hypothesis when it is true is called an error of the first kind, and the choice of  $\alpha$  therefore limits the risk of such an error.

On the other hand, it is possible to commit an error of the second kind, that is, accepting the null hypothesis when it is false. The probability  $1 - \beta$  of rejecting the null hypothesis when it is false is called the power of the test; the probability of an error of the second kind is therefore  $\beta$ .

For a given sample *n* and error of the first kind, these probabilities depend not only on the true mean *D* of the observed differences  $d_i = X_i - Y_i$  for which one can postulate different alternative hypotheses but also on the standard deviation  $\sigma_d$  of these differences. This standard deviation is in general unknown and if *n* is small the sample will provide only a poor estimator.

The result is that it is impossible to set an upper limit to the probability of an error of the second kind. Nevertheless, in the following graphs the relation is shown between the power of the test,  $1 - \beta$ , and the actual population mean divided by the corresponding standard deviation,  $D/\sigma_d$  for one-sided tests of the hypothesis  $H_0: D \leq 0$ , and for various values of n and for the significance levels 0,05 and 0,01 respectively.

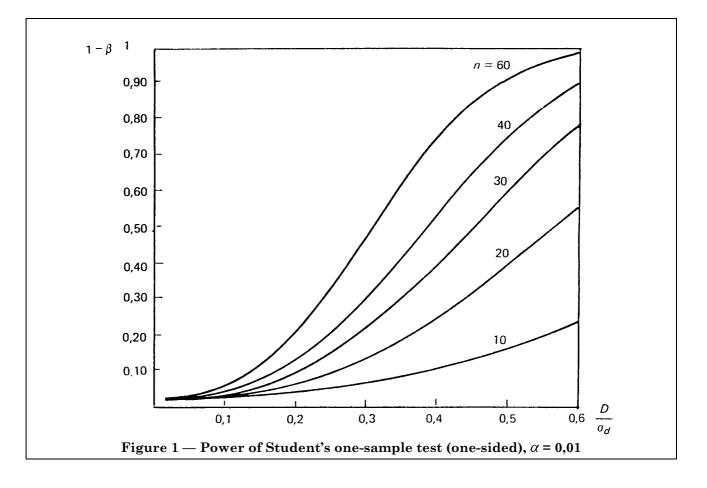
From these graphs the following conclusions may be drawn:

1) The power of the test is uniquely determined by the true mean of the differences, measured in units of their standard deviation, by the significance level  $\alpha$  and the sample size.

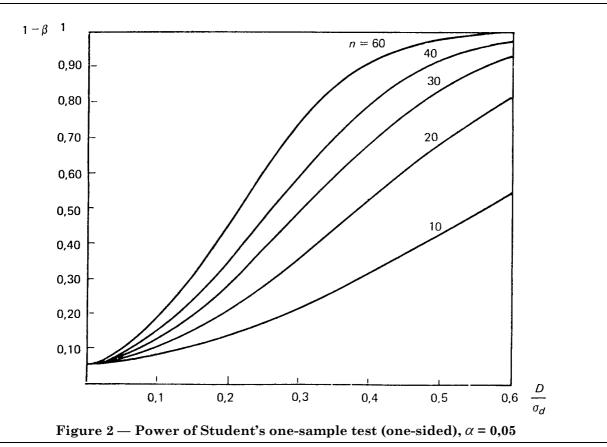
2) The power function is a strictly increasing function of the true mean difference.

It is also strictly increasing with the sample size and the significance level  $\alpha$ , provided D > 0and  $\alpha$  different from 0 and from 1.

3) With a significance level of 0,05 and a sample size of 50, a power of at least 0,95 is already obtained when the true mean difference exceeds one-half of the standard deviation of the differences. For n = 20, this power is obtained for  $D/\sigma_d = 0,78$  or more.



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NOTE The graphs are based on the work of D.B. OWEN, Handbook of statistical tables, Addison Wesley.

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